- It was suggested first in [1] that nothing really imposed that the vectors of the reconstructed attractor should be composed of equally spaced samples. Actually, what matters is the time span (*m*-1)τ covered by the vectors.
- But the approach proposed consists in selecting terms in a linear AR model and retaining those as vector components in the embedding.





1

• In [2] this idea was somewhat amplified. Let us consider now embedded vectors

$$X_{n-1} = [x(n-i_1), x(n-i_2), \dots, x(n-i_{\max})]^{\mathsf{T}}$$

with all  $i_k < d$ , d chosen beforehand, and one tries to predict sample x(n) using some  $G[X_{n-1}]$ .

- Selection of the indices amounts to selecting a bit string *a* of length *d*.
- The idea is to select the best possible *G* and *a* using an MDL approach.





- Recall that the coding length of data + model is:  $L(x,\theta,a) = -\log P(x|\theta,a) + L(\theta) + L(a)$ where  $\theta$  is the model parameter vector.
- The log likelihood of the data  $\{x(n)\}, n = 1, ..., N$ , must include the log likelihood of the initial conditions  $X^{(0)} = [x(d), x(2), ..., x(1)]^{\mathsf{T}}$ :
  - $-\log P(x|\theta,a) = -\log P(x|\theta,a,X^{(0)}) \log P(X^{(0)})$





- If it is assumed that the errors and the initial conditions are independent and normally distributed, one gets:
  - $L(x,\theta,a) = -\log P(x|\mathsf{N}(0,\sigma^2)) \log P(X^{(0)}|\mathsf{N}(0,\sigma_X^2)) + L(\theta) + L(\theta)$
- Now, in addition to determining the best set of indices, one must also select the best model.





- The idea proposed in [2] consists in using the deterministic prediction scheme as the model. Any imbedded vector can be used for prediction.
- This approach is simple to implement, and it is robust.
- It does not require any parameter, thus  $L(\theta) = 0$ .
- But it is only a pseudo-model, since prediction of x(n) may imply  $X_k$ , with k > n.





• After some manipulations, and suppression of constant or negligible terms, the MDL criterion is:

$$MDL(a) \approx \frac{d}{2} \log \left[ \frac{1}{d} \sum_{i=1}^{d} (x(i) - \overline{x})^{2} \right] \xleftarrow{\text{initial}}_{\text{conditions}}$$
$$+ \frac{N - d}{2} \log \left[ \frac{1}{N - d} \sum_{i=d+1}^{N} e(i)^{2} \right] + d$$
$$\operatornamewithlimits{\uparrow}_{\text{errors}} \det{\text{ordel}}$$





- Examining all possible combinations of indices up to a maximum value *d* means estimating 2<sup>d</sup> models. This is not feasible.
- What is proposed in [2] is a sequential selection procedure, i.e.:
- the first term is the one giving the smallest MDL and kept.
- All possible 2nd terms are tested and the best one kept if the MDL is smaller ...





- Unfortunately, this sequential selection does not work as well as for term selection in pseudolinear models, which amounts to defining the best subspace to project to.
- Why not apply a genetic algorithm? Each possible embedding is coded as a bit string, with a '1' if the component is selected, '0' otherwise.





• Lorenz system, *x* coordinate

Sequential and GA: Same model Indices 1, 2 (maximum 15)







• Sunspot time series

Sequential Indices 1, 2, 5 MDL: 839.02

#### GA Indices 1, 2, 3, 5, 11 MDL: 837.5

200 150 100 50 -50 50 100 150 200 250 300 'n Blue: signal. Red: prediction error 200 150 100 50 Π -50 -100 50 150 200 250 300

Blue: signal. Red: prediction error

# (maximum 15)





• Surrogate on sunspot time series

# Sequential Indices 1, 3, 12, 13 MDL: 931.62

#### GA Indices 1, 3, 12, 13 MDL: 931.62



### (maximum 15)





• RR Intervals (between heartbeats), sampling 2 Hz.







• Increase in model size with noise (Henon map, 300 samples)

snr (dB)	terms
20	1, 2, 3
15	1, 2, 3
10	1, 2, 3, 8
5	1, 2, 3, 4, 7, 8, 12, 15
0	1, 2, 3, 5, 7, 9, 10, 11, 13, 14





- A technique to assess the interdependence between two multivariate time series, related to deterministic prediction, has been proposed [3,4].
- Of course, those multivariate time series X = {X<sub>n</sub>} and Y = {Y<sub>n</sub>} may result from the embedding of two signals {x(k)} and {y(k)}.
- The idea is to look for the possible existence of (well behaved) functions *G*(.) and *F*(.) such that

$$X_n = G(Y_n)$$
 and  $Y_n = F(X_n)$ 





- It is to be noted that interdependence does not imply causation. The relationship between the two series may be due to the influence of a third one on both of them.
- The principle of the test is simple and robust: if indeed G(.) and F(.) exist and are reasonably continuous, then if  $Y_n$  is close to  $Y_k$ , then  $X_n = G(Y_n)$  should be close to  $X_k = G(Y_k)$ .
- Closeness is quantified with respect to a deterministic prediction of  $X_n$  using closest neighbors in  $\{X_k\}$ .





- Let  $r_{n,j}$  and  $s_{n,j}$ , j = 1, ..., k, denote the indices of the *k* nearest neighbors of  $X_n$  and  $Y_n$  respectively.
- The mean squared distance between  $X_n$  and its closest neighbors is:

$$R_n^{(k)}(\mathbf{X}) = \frac{1}{k} \sum_{j=1}^k ||X_n - X_{r_{n,j}}||^2$$

This quantity should be small for close enough neighbors.





- The *conditional* mean squared distance with respect to  $\{Y_k\}$  is:  $R_n^{(k)}(\mathbf{X}|\mathbf{Y}) = \frac{1}{k} \sum_{i=1}^k ||X_n - X_{s_{n,j}}||^2$
- One can define the same quantities for  $Y_n$ :

$$R_n^{(k)}(\mathbf{Y}) = \frac{1}{k} \sum_{j=1}^{k} ||Y_n - Y_{s_{n,j}}||^2 \qquad R_n^{(k)}(\mathbf{Y}|\mathbf{X}) = \frac{1}{k} \sum_{j=1}^{k} ||Y_n - Y_{r_{n,j}}||^2$$



If the point set {Y<sub>k</sub>} has an average squared radius
 R(X):

$$R_n^{(k)}(\mathbf{X})/R(\mathbf{X}) \approx (k/N)^{2/m} <<1 \text{ for } k << N$$

with *m* the dimension and *N* the number of vectors.

• If X and Y are indeed interdependent, then the same relationship should hold for the conditional mean square distance.





• A measure of local interdependence is:

$$S_n^{(k)}(\mathbf{X} | \mathbf{Y}) = \frac{R_n^{(k)}(\mathbf{X})}{R_n^{(k)}(\mathbf{X} | \mathbf{Y})} \qquad 0 < S_n^{(k)}(\mathbf{X}) \le 1$$

• And the global measure is:

$$S^{(k)}(\mathbf{X}|\mathbf{Y}) = \frac{1}{N} \sum_{n=1}^{N} S_n^{(k)}(\mathbf{X}|\mathbf{Y}) \qquad 0 < S^{(k)}(\mathbf{X}|\mathbf{Y}) \le 1$$

• This global measure should be close to 0 when there is no interdependence, and close to 1 for strong interdependence





- To assess the significance of S<sup>(k)</sup>(X|Y), one can generate several surrogate data Y<sup>(u)</sup> for Y, compute S<sup>(k)</sup>(X|Y<sup>(u)</sup>), and perform a rank test.
- If there is indeed a significant interdependence, it is possible to test if it is linked to phase only by generating bivariate surrogated data {X<sup>(b)</sup>, Y<sup>(b)</sup>}, compute S<sup>(k)</sup>(X<sup>(b)</sup>|Y<sup>(b)</sup>), and perform a rank test.
- Those bivariate surrogate data are generated by applying the same phase randomization to both signals.

![](_page_19_Picture_4.jpeg)

![](_page_19_Picture_6.jpeg)

- Recurrence plots (RP) were initially introduced to display the recurrence of patterns and possible non stationarities in imbedded time series [5,6].
- Apart from producing nice pictures, several parameters can be extracted from RP to characterize dynamical processes.
- RP have also been generalized to the analysis of the interdependence between two time series.

![](_page_20_Picture_4.jpeg)

![](_page_20_Picture_6.jpeg)

- The principle of RP is quite simple: given a multivariate time series {X<sub>n</sub>} n = 1, ..., N, possibly obtained through the embedding of a signal, an RP is an N×N array in which a dot is placed at location (*i*,*j*) if ||X<sub>i</sub> X<sub>j</sub>|| < δ, δ a predetermined small number.</li>
- Typically,  $\delta$  is a fraction of the sum of standard deviations of the vector components (sstd).
- Of course, RPs are symmetrical with respect to the main diagonal.

![](_page_21_Picture_4.jpeg)

![](_page_21_Picture_6.jpeg)

- Periodicities (recurrences) are expressed by diagonal lines in RPs. RP-based analysis is mostly based on the characterization of these lines.
- Of course, an RP is dependent upon the threshold
  δ, but it is also quite dependent upon vector
  dimension *m*.
- What matters is to conserve these values through the various experiments performed.

![](_page_22_Picture_4.jpeg)

![](_page_22_Picture_6.jpeg)

• Example: sinusoid with period 20, m = 1,  $\delta = 0.05$  sstd.

![](_page_23_Figure_2.jpeg)

![](_page_23_Picture_3.jpeg)

![](_page_23_Picture_4.jpeg)

• Example: sinusoid with period 20, m = 2,  $\delta = 0.05$  sstd.

![](_page_24_Figure_2.jpeg)

![](_page_24_Picture_3.jpeg)

![](_page_24_Picture_4.jpeg)

• Example: Lorenz,  $[x \ y \ z]$ ,  $\delta = 0.05$  sstd.

![](_page_25_Picture_2.jpeg)

![](_page_25_Picture_3.jpeg)

![](_page_25_Picture_4.jpeg)

![](_page_25_Picture_5.jpeg)

Parameters extracted:

- REC proportion of recurrence (black) points. Higher for periodic dynamics.
- DET proportion of recurrence points in diagonal lines of length at least 2. Higher for deterministic dynamics.
- ENT entropy of line lengths on the diagonal. Higher for more complex dynamics

![](_page_26_Picture_5.jpeg)

![](_page_26_Picture_7.jpeg)

- DIV inverse of the length of the longest diagonal line. Higher for larger maximum Lyapunov exponent.
- TREND slope *a* of the regression line:

$$R_j = b + ad_j + \mathcal{E}_j$$

with  $R_j$  the proportion of recurrence points on diagonal at distance  $d_j$  from the main diagonal. Quantifies the decrease in recurrence with respect to time difference.

![](_page_27_Picture_5.jpeg)

![](_page_27_Picture_7.jpeg)

• Clearly, significance of the values of these parameters in terms of determinism or nonlinearity must be assessed using surrogate data.

![](_page_28_Picture_2.jpeg)

![](_page_28_Picture_3.jpeg)

REC=0.03 DET=0.86 DIV=0.09 trend=-1.5e-5

![](_page_28_Figure_5.jpeg)

REC=0.008 DET=0.24 DIV=0.33 trend=-1.7e-5

![](_page_28_Picture_7.jpeg)

![](_page_28_Picture_9.jpeg)

- It is also possible to generate RP with two time series (same range, same dimension).
- Example: x and y of Lorenz

![](_page_29_Picture_4.jpeg)

LTS

![](_page_29_Picture_6.jpeg)

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![](_page_30_Picture_7.jpeg)

![](_page_30_Picture_9.jpeg)