

SOLUTION 1.

- (a) At first look it may seem that the probability is uniformly distributed over the disk, but in the next part we will show that this is not true.
- (b) We know that  $R$  is uniformly distributed in  $[0, 1]$  and  $\Phi$  is uniformly distributed in  $[0, 2\pi)$ , so we have  $f_R(r) = 1$  if  $0 \leq r \leq 1$  and  $f_\Phi(\phi) = \frac{1}{2\pi}$  if  $0 \leq \phi < 2\pi$ .

As these two random variables are independent, we have

$$f_{R,\Phi}(r, \phi) = \begin{cases} \frac{1}{2\pi} & 0 \leq r \leq 1 \text{ and } 0 \leq \phi < 2\pi \\ 0 & \text{otherwise.} \end{cases}$$

It can be easily shown that the Jacobian determinant is  $\det J = r = \sqrt{x^2 + y^2}$ . Therefore, the probability distribution in cartesian coordinates is

$$\begin{aligned} f_{X,Y}(x, y) &= \frac{1}{|\det J|} f_{R,\Phi}(r, \phi) \\ &= \begin{cases} \frac{1}{2\pi\sqrt{x^2+y^2}} & x^2 + y^2 \leq 1 \\ 0 & \text{otherwise.} \end{cases} \end{aligned}$$

- (c) We see that the probability distribution is not distributed uniformly. This makes sense because rings of equal width have the same probability but not the same area.

SOLUTION 2.

- (a) Let the two hypotheses be  $H = 0$  and  $H = 1$  when  $c_0$  and  $c_1$  are transmitted, respectively. The ML decision rule is

$$f_{Y_1 Y_2 | H}(y_1, y_2 | 1) \underset{\hat{H}=0}{\overset{\hat{H}=1}{\gtrless}} f_{Y_1 Y_2 | H}(y_1, y_2 | 0).$$

Because  $Z_1$  and  $Z_2$  are independent, we can write

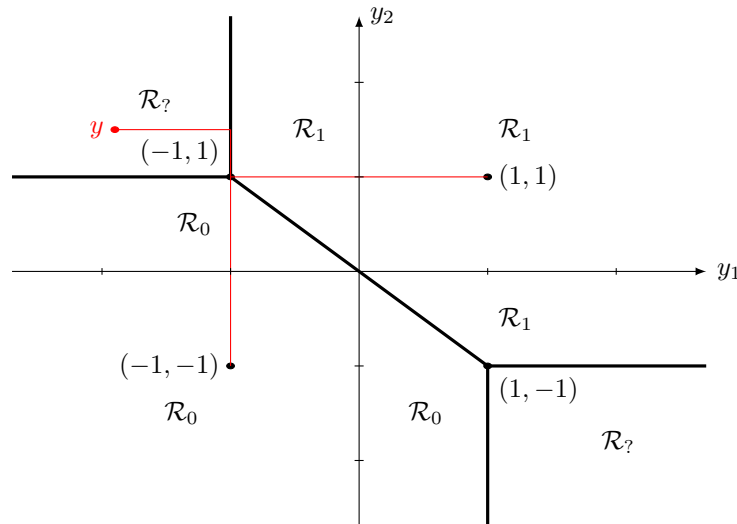
$$\frac{1}{2} e^{-|y_1-1|} \frac{1}{2} e^{-|y_2-1|} \underset{\hat{H}=0}{\overset{\hat{H}=1}{\gtrless}} \frac{1}{2} e^{-|y_1+1|} \frac{1}{2} e^{-|y_2+1|},$$

and, after taking the logarithm,

$$|y_1 + 1| + |y_2 + 1| \underset{\hat{H}=0}{\overset{\hat{H}=1}{\gtrless}} |y_1 - 1| + |y_2 - 1|.$$

- (b) Because the hypotheses are equally likely and  $Z_1$  and  $Z_2$  have the same distribution, the decision region for  $\hat{H} = 0$  contains the points closer to  $(-1, -1)$  and the decision region for  $\hat{H} = 1$  contains the points closer to  $(1, 1)$ . For this problem, the distance between the points  $(y_{11}, y_{12})$  and  $(y_{21}, y_{22})$  is the Manhattan distance,  $|y_{11} - y_{21}| + |y_{12} - y_{22}|$ , and not the Euclidian distance.

Let us first consider the points above the line  $y_2 = -y_1$  in the figure below. It is easy to notice that the points in the positive quadrant are closer to  $(1, 1)$  than to  $(-1, -1)$ , therefore they belong to  $\mathcal{R}_1$  ( $\hat{H} = 1$ ). This is also true if  $\{(y_1 \geq 0) \cap (y_2 \in (-1, 0))\}$ , or if  $\{(y_2 \geq 0) \cap (y_1 \in (-1, 0))\}$ .



Similar reasoning can be applied to the points below the diagonal to determine  $\mathcal{R}_0$ .

The points for which  $\{(y_1 \leq -1) \cap (y_2 \geq 1)\}$  or  $\{(y_1 \geq 1) \cap (y_2 \leq -1)\}$  are equally distanced to  $(-1, -1)$  and  $(1, 1)$ , therefore they can belong to either  $\mathcal{R}_0$  or  $\mathcal{R}_1$  with the same probability. This region is named  $\mathcal{R}_2$ .

- (c) The two hypotheses are equally probable for the region  $\mathcal{R}_2$ . Therefore, we can split this region in any way between the decision regions and have the same error probability. Because  $\mathcal{R}_1$  is included in the region for which  $y_2 > -y_1$  and  $\mathcal{R}_0$  does not intersect the region for which  $y_2 > -y_1$ , the error probability is minimized by deciding  $\hat{H} = 1$  if  $(y_1 + y_2) > 0$ .

- (d)

$$\begin{aligned}
 P_e(0) &= \Pr\{Y_1 + Y_2 > 0 | H = 0\} \\
 &= \Pr\{Z_1 + Z_2 - 2 > 0\} \\
 &= \int_2^\infty \frac{e^{-w}}{4} (1 + w) dw \\
 &= \frac{-e^{-w}}{4} (w + 2) \Big|_2^\infty = e^{-2}.
 \end{aligned}$$

By symmetry, and considering that the messages are equally likely,  $P_e(0) = P_e(1) = P_e$ .

SOLUTION 3. The first basis vector is the first waveform after normalization. We first compute  $\|w_0(t)\|$ .

$$\|w_0(t)\| = \sqrt{\int |w_0(t)|^2 dt} = \sqrt{\int_0^T 1 dt} = \sqrt{T}$$

$$\psi_0(t) = \frac{w_0(t)}{\|w_0(t)\|} = \frac{w_0(t)}{\sqrt{T}} = \begin{cases} \frac{1}{\sqrt{T}} & \text{if } 0 \leq t \leq T \\ 0 & \text{otherwise} \end{cases}$$

We get the second basis vector as follows:

$$\langle w_1(t), \psi_0(t) \rangle = \int_0^{\frac{T}{2}} \frac{2}{\sqrt{T}} dt = \sqrt{T}$$

$$\alpha_1(t) = w_1(t) - \langle w_1(t), \psi_0(t) \rangle \psi_0(t) = w_1(t) - w_0(t) = \begin{cases} 1 & \text{if } 0 \leq t \leq \frac{T}{2} \\ -1 & \text{if } \frac{T}{2} < t \leq T \\ 0 & \text{otherwise} \end{cases}$$

$$\psi_1(t) = \frac{\alpha_1(t)}{\|\alpha_1(t)\|} = \begin{cases} \frac{1}{\sqrt{T}} & \text{if } 0 \leq t \leq \frac{T}{2} \\ -\frac{1}{\sqrt{T}} & \text{if } \frac{T}{2} < t \leq T \\ 0 & \text{otherwise} \end{cases}$$

SOLUTION 4.

(a) We use the Gram-Schmidt procedure:

- 1) The first step is to normalize the function  $\beta_0(t)$ , i.e. the first function of the basis that we are looking for is

$$\begin{aligned} \psi_0(t) &= \frac{\beta_0(t)}{\|\beta_0(t)\|} = \frac{\beta_0(t)}{\sqrt{\int \beta_0(t)^2 dt}} \\ &= \frac{\beta_0(t)}{\sqrt{\int_0^1 4t^2 dt}} = \frac{\sqrt{3}}{2} \beta_0(t) = \begin{cases} 0 & \text{if } t < 0 \\ \sqrt{3}t & \text{if } 0 \leq t \leq 1 \\ 0 & \text{if } t > 1 \end{cases} \end{aligned}$$

- 2) Next, we subtract from  $\beta_1(t)$  the components that are in the span of the currently established part of the basis, i.e. in the span of  $\{\psi_0(t)\}$ . This can be achieved by projecting  $\beta_1(t)$  onto  $\psi_0(t)$  and then subtracting this projection from  $\beta_1(t)$ , i.e.

$$\begin{aligned} \alpha_1(t) &= \beta_1(t) - \langle \beta_1(t), \psi_0(t) \rangle \psi_0(t) = \beta_1(t) - \left( \int \beta_1(t) \psi_0(t) dt \right) \psi_0(t) \\ &= \beta_1(t) - \left( \frac{\sqrt{3}}{2} \right) \left( \frac{4}{3} \right) \psi_0(t) \\ &= \beta_1(t) - \frac{2}{\sqrt{3}} \psi_0(t) \\ &= \beta_1(t) - \beta_0(t). \end{aligned}$$

From this, we find the second basis element as

$$\psi_1(t) = \frac{\alpha_1(t)}{\|\alpha_1(t)\|} = \begin{cases} 0 & \text{if } t < 1 \\ -\sqrt{3}(t-2) & \text{if } 1 \leq t \leq 2 \\ 0 & \text{if } t > 2 \end{cases}$$

3) Again, we subtract from  $\beta_2(t)$  the components that are in the span of the currently established part of the basis, i.e. in the span of  $\{\psi_0(t), \psi_1(t)\}$ . This can be achieved by projecting  $\beta_2(t)$  onto  $\psi_0(t)$  and  $\psi_1(t)$  and then subtracting both these projections from  $\beta_2(t)$ . For this step, it is *essential* that the basis elements  $\{\psi_0(t), \psi_1(t)\}$  be orthonormal. Continuing the derivation, we obtain

$$\begin{aligned}\alpha_2(t) &= \beta_2(t) - \langle \beta_2(t), \psi_0(t) \rangle \psi_0(t) - \langle \beta_2(t), \psi_1(t) \rangle \psi_1(t) \\ &= \beta_2(t) - \left( \int \beta_2(t) \psi_0(t) dt \right) \psi_0(t) - \left( \int \beta_2(t) \psi_1(t) dt \right) \psi_1(t) \\ &= \beta_2(t) - 0 - \alpha_1(t) \\ &= \beta_2(t) - \beta_0(t) + \beta_1(t),\end{aligned}$$

and from this, we find the third basis element as

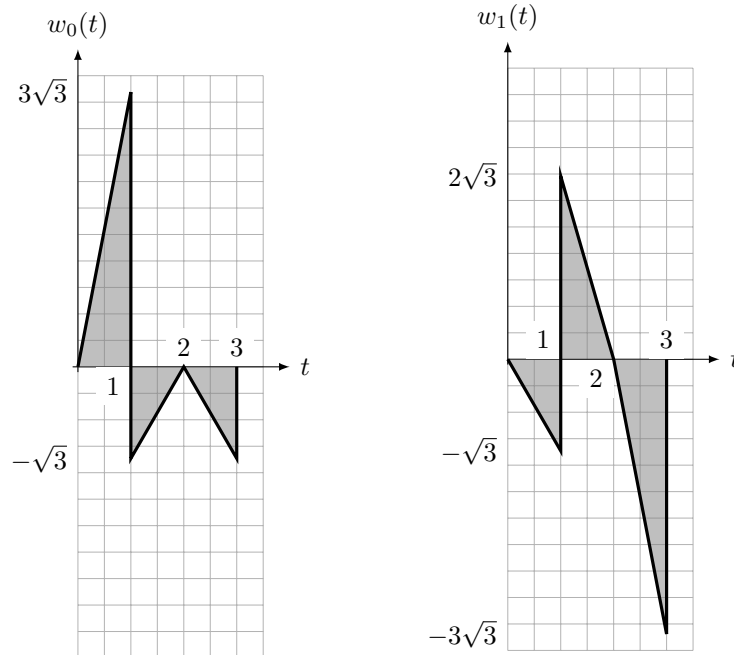
$$\psi_2(t) = \frac{\alpha_2(t)}{\|\alpha_2(t)\|} = \begin{cases} 0 & \text{if } t < 2 \\ -\sqrt{3}(t-2) & \text{if } 2 \leq t \leq 3 \\ 0 & \text{if } t > 3 \end{cases}$$

(b) By definition we can write  $w_0(t)$  and  $w_1(t)$  as follows

$$w_0(t) = 3\psi_0(t) - \psi_1(t) + \psi_2(t) = \begin{cases} 3\sqrt{3}t & \text{if } 0 \leq t < 1 \\ \sqrt{3}(t-2) & \text{if } 1 < t < 2 \\ -\sqrt{3}(t-2) & \text{if } 2 < t \leq 3 \end{cases}$$

and

$$w_1(t) = -\psi_0(t) + 2\psi_1(t) + 3\psi_2(t) = \begin{cases} -\sqrt{3}t & \text{if } 0 \leq t < 1 \\ -2\sqrt{3}(t-2) & \text{if } 1 < t < 2 \\ -3\sqrt{3}(t-2) & \text{if } 2 < t \leq 3 \end{cases}$$



(c)

$$\langle c_0, c_1 \rangle = -3 \cdot 1 - 1 \cdot 2 + 1 \cdot 3 = -2.$$

We know that  $w_0(t)$  and  $w_1(t)$  are both real, thus

$$\begin{aligned} \langle w_0(t), w_1(t) \rangle &= \int w_0(t)w_1(t) dt = \int_0^1 -9t^2 dt + \int_1^2 -6(t-2)^2 dt + \int_2^3 9(t-2)^2 dt \\ &= - \int_1^2 6(t-2)^2 dt = -2. \end{aligned}$$

We see that the inner products are equal as expected.

(d)

$$\begin{aligned} \|c_0\| &= \sqrt{\langle c_0, c_0 \rangle} = \sqrt{11}, \\ \|w_0\|^2 &= \int |w_0(t)|^2 dt = \int_0^1 27t^2 dt + \int_1^3 3(t-2)^2 dt = 9 + 2 = 11. \end{aligned}$$

We see that the norms are also equal.

SOLUTION 5.

(a)

$$\|g_i\| = \sqrt{T}, \quad i = 1, 2, 3.$$

(b)  $Z_1$  and  $Z_2$  are independent since  $g_1$  and  $g_2$  are orthogonal. Hence  $Z$  is a Gaussian random vector  $\sim \mathcal{N}(0, \sigma^2 I_2)$ , where  $\sigma^2 = \frac{N_0}{2}T$ .

(c)

$$\begin{aligned} P_a &= \Pr\{Z_1 \in [1, 2] \cap Z_2 \in [1, 2]\} = \Pr\{Z_1 \in [1, 2]\} \Pr\{Z_2 \in [1, 2]\} \\ &= \left[ Q\left(\frac{1}{\sigma}\right) - Q\left(\frac{2}{\sigma}\right) \right]^2, \end{aligned}$$

where  $\sigma^2 = \frac{N_0}{2}T$ .

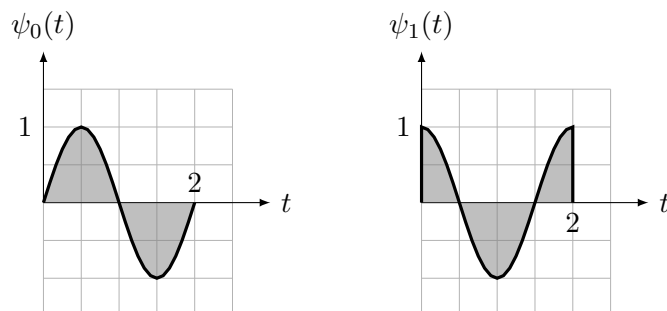
(d)  $P_b = P_a$ , since one obtains the square (b) from the square (a) via a rotation.

(e)  $Z_3 = -Z_1$ .  $U = Z_1(1, -1)^T$ , and thus  $U$  can never be in (a), hence  $Q_a = 0$ .

(f)  $U$  is in square (c) if and only if  $Z_1 \in [1, 2]$ . Hence  $Q_c = Q\left(\frac{1}{\sigma}\right) - Q\left(\frac{2}{\sigma}\right)$ , where  $\sigma^2 = \frac{N_0}{2}T$ .

SOLUTION 6.

(a) An orthonormal basis for the signal space spanned by the waveforms is<sup>1</sup>:



<sup>1</sup>this can be obtained using the Gram-Schmidt procedure or simply by looking at the waveforms.

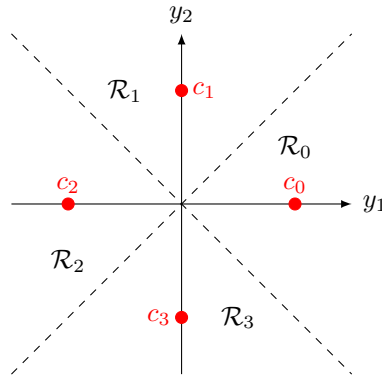
(b) The codewords representing the waveforms are

$$\begin{aligned} c_0 &= (\sqrt{\mathcal{E}}, 0) \\ c_1 &= (0, \sqrt{\mathcal{E}}) \\ c_2 &= (-\sqrt{\mathcal{E}}, 0) \\ c_3 &= (0, -\sqrt{\mathcal{E}}) \end{aligned}$$

(c) As we have seen in the lecture, if  $R(t)$  is the noisy received waveform,  $(Y_0, Y_1) = (\langle R, \psi_0 \rangle, \langle R, \psi_1 \rangle)$  is a sufficient statistic for decision. Hence, we have the following hypothesis testing problem: Under  $H = i, i = 0, 1, 2, 3$ ,

$$Y_i = c_i + Z,$$

where  $Z \sim \mathcal{N}(0, \frac{N_0}{2} I_2)$ . One can check that  $c_i, i = 0, 1, 2, 3$  represent the QPSK codewords, and the decision regions for the ML receiver will be as follows:



The distance between two adjacent codewords (say  $c_0$  and  $c_1$ ) is  $d = \sqrt{2\mathcal{E}}$  and the error probability of the receiver is

$$\begin{aligned} P_e &= 2Q\left(\frac{d}{2\sigma}\right) - Q^2\left(\frac{d}{2\sigma}\right) \\ &= 2Q\left(\frac{\sqrt{2\mathcal{E}}}{2\sqrt{N_0/2}}\right) - Q^2\left(\frac{\sqrt{2\mathcal{E}}}{2\sqrt{N_0/2}}\right) \\ &= 2Q\left(\sqrt{\frac{\mathcal{E}}{N_0}}\right) - Q^2\left(\sqrt{\frac{\mathcal{E}}{N_0}}\right). \end{aligned}$$