

ÉCOLE POLYTECHNIQUE FÉDÉRALE DE LAUSANNE

School of Computer and Communication Sciences

Handout 3

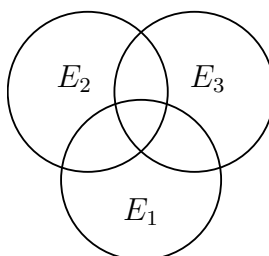
Solutions to homework 1

Information Theory and Coding

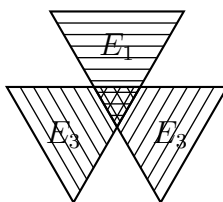
Sep. 25, 2017

PROBLEM 1. Note that $E_0 = E_1 \cup E_2 \cup E_3$.

- (a) (1) For disjoint events, $P(E_0) = P(E_1) + P(E_2) + P(E_3)$, so $P(E_0) = 3/4$.
- (2) For independent events, $1 - P(E_0)$ is the probability that none of the events occur, which is the product of the probabilities that each one doesn't occur. Thus $1 - P(E_0) = (3/4)^3$ and $P(E_0) = 37/64$.
- (3) If $E_1 = E_2 = E_3$, then $E_0 = E_1$ and $P(E_0) = 1/4$.
- (b) (1) From the following Venn diagram, $P(E_0)$ is clearly maximized when the events are disjoint, so $\max P(E_0) = 3/4$.



- (2) The intersection of each pair of sets has probability $1/16$. As seen in the Venn diagram below, $P(E_0)$ is maximized if all these pairwise intersections are identical, in which case $P(E_0) = 3(1/4 - 1/16) + 1/16 = 5/8$. One can also use the formula $P(E_0) = P(E_1) + P(E_2) + P(E_3) - P(E_1 \cap E_2) - P(E_1 \cap E_3) - P(E_2 \cap E_3) + P(E_1 \cap E_2 \cap E_3)$, and notice that all the terms except the last is fixed by the problem, and the last term cannot be made more than $\min_{i,j} P(E_i \cap E_j) = 1/16$.



- (c) Same considerations as in (b)(2) yields the upper bound $P(E_0) \leq 3p - 2p^2$. As $P(E_0) = 1$, we find that $p \geq 1/2$.

PROBLEM 2. Let L be the event that the loaded die is picked and H the event that the honest die is picked. Let A_i be the event that i is turned up on the first roll, and B_i be the event that i is turned up on the second roll. We are given that $P(L) = 1/3, P(H) = 2/3; P(A_1 | L) = 2/3; P(A_i | L) = 1/15 \quad 2 \leq i \leq 6; P(A_i | H) = 1/6 \quad 1 \leq i \leq 6$. Then

$$P(L | A_1) = \frac{P(L, A_1)}{P(A_1)} = \frac{P(A_1 | L) P(L)}{P(A_1 | L) P(L) + P(A_1 | H) P(H)} = \frac{2}{3}.$$

This is the probability that the loaded die was picked conditional on the first roll showing a 1. For two rolls we make the assumption from the physical mechanism involved in rolling a die that the outcome on the two successive rolls of a given die are independent. Thus $P(A_1B_1 | L) = (2/3)^2$ and $P(A_1B_1 | H) = (1/6)^2$. It follows as before that

$$P(L | A_1B_1) = \frac{8}{9}.$$

PROBLEM 3. Since A, B, C, D form a Markov chain their probability distribution is given by

$$p(a)p(b|a)p(c|b)p(d|c) \tag{1}$$

(a) Yes: Summing (1) over d shows that A, B, C have the probability distribution $p(a)p(b|a)p(c|b)$.

(b) Yes: The reverse of a Markov chain is also a Markov chain. Applying this to A, B, C, D and using part (a) we get that D, C, B is a Markov chain. Reversing again we get the desired result.

Alternatively summing (1) over a (noting that $\sum_a p(a)p(b|a) = p(b)$) shows $B \oplus C \oplus D$.

(c) Yes: Since A, B, C, D is a Markov chain, given C, D is independent of B , and thus $p(d|c) = p(d|(b, c))$. So (1) can be written as

$$p(a, (b, c), d) = p(a)p((b, c)|a)p(d|(b, c)).$$

(d) Yes, by a similar (in fact easier) reasoning as (c).

PROBLEM 4. No. Take for example $A = D$ and let A be independent of the pair (B, C) . Then both A, B, C and B, C, A (same as B, C, D) are Markov chains. But A, B, C, D is not: A is not independent of D when B and C are given.

PROBLEM 5.

(a)

$$\begin{aligned} E[X + Y] &= \sum_{x,y} (x + y)P_{XY}(x, y) \\ &= \sum_{x,y} xP_{XY}(x, y) + \sum_{x,y} yP_{XY}(x, y) \\ &= \sum_x xP_X(x) + \sum_y yP_Y(y) \\ &= E[X] + E[Y]. \end{aligned}$$

Note that independence is not necessary here and that the argument extends to non-discrete variables if the expectation exists.

(b)

$$\begin{aligned} E[XY] &= \sum_{x,y} xyP_{XY}(x, y) \\ &= \sum_{x,y} xyP_X(x)P_Y(y) \\ &= \sum_x xP_X(x) \sum_y yP_Y(y) \\ &= E[X] E[Y]. \end{aligned}$$

Note that the statistical independence was used on the second line. Let X and Y take on only the values ± 1 and 0 . An example of uncorrelated but dependent variables is

$$P_{XY}(1, 0) = P_{XY}(0, 1) = P_{XY}(-1, 0) = P_{XY}(0, -1) = \frac{1}{4}.$$

An example of correlated and dependent variables is

$$P_{XY}(1, 1) = P_{XY}(-1, -1) = \frac{1}{2}.$$

(c) Using (a), we have

$$\begin{aligned}\sigma_{X+Y}^2 &= E[(X - E[X] + Y - E[Y])^2] \\ &= E[(X - E[X])^2] + 2E[(X - E[X])(Y - E[Y])] + E[(Y - E[Y])^2].\end{aligned}$$

The middle term, from (a), is $2(E[XY] - E[X]E[Y])$. For uncorrelated variables that is zero, leaving us with $\sigma_{X+Y}^2 = \sigma_X^2 + \sigma_Y^2$.

PROBLEM 6. We solve the problem for a general vehicle with n wheels.

- (a) Out of $n!$ possible orderings $(n - 1)!$ has the tyre 1 in its original place. Thus tyre 1 is installed in its original position with probability $1/n$.
- (b) All tyres end up in their original position in only 1 of the $n!$ orders. Thus the probability of this event is $1/n!$.
- (c) Let X_i be the indicator random variable that tyre i is installed in its original position, so that the number of tyres installed in their original positions is $N = \sum_{i=1}^n X_i$. By (a), $E[X_i] = 1/n$. By the linearity of expectation, $E[N] = n(1/n) = 1$. Note that the linearity of the expectation holds even if the X_i 's are not independent (as it is in this case).
- (e) Let A_i be the event that the i th tyre remains in its original position. Then, the event we are interested in is the complement of the event $\bigcup_i A_i$ and thus has probability $1 - \Pr(\bigcup_i A_i)$. Furthermore, by the inclusion/exclusion formula,

$$\Pr\left(\bigcup_i A_i\right) = \sum_i \Pr(A_i) - \sum_{i_1 < i_2} \Pr(A_{i_1} \cap A_{i_2}) + \sum_{i_1 < i_2 < i_3} \Pr(A_{i_1} \cap A_{i_2} \cap A_{i_3}) - \dots$$

The j th sum above consists of $\binom{n}{j}$ terms, each term having the value $\Pr(A_1 \cap \dots \cap A_j)$. Note that this is the probability of the event that tyres 1 through j have remained in their original positions, and equals $(n - j)!/n!$. Consequently,

$$\Pr\left(\bigcup_i A_i\right) = \sum_{j=1}^n (-1)^{j-1} \binom{n}{j} \frac{(n - j)!}{n!} = \sum_{j=1}^n (-1)^{j-1} 1/j!,$$

and the event that no tyre remains in its original position has probability

$$1 - \Pr\left(\bigcup_i A_i\right) = \sum_{j=0}^n \frac{(-1)^j}{j!}.$$

(For the case $n = 4$, the value is $3/8$.)

PROBLEM 7.

- (a) Let A_i denote the event that $X_i \neq X$. The event that X does not appear in the inventory is thus

$$A = A_1 \cap \cdots \cap A_n.$$

Note that the events A_1, \dots, A_n are not independent—because they involve the common random variable X . However, they become independent when conditioned on the value of X , with $P(A_i|X = x) = 1 - p(x)$. Thus,

$$P(A|X = x) = (1 - p(x))^n.$$

Consequently $P(A) = \sum_x p(x)(1 - p(x))^n$.

- (b) With p the uniform distribution on n items, the above value for $P(A)$ equals $(1 - 1/n)^n$.
- (c) For n large, $(1 - 1/n)^n$ approaches $1/e \approx 37\%$.